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# THE PREVALENCE AND EFFECT ON HOURS WORKED OF THE MINIMUM WAGE IN IRELAND

A SECTORAL AND REGIONAL ANALYSIS

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This report has been accepted for publication by the Institute, which does not itself take institutional policy positions. All ESRI Research Series reports are peer reviewed prior to publication. The authors are solely responsible for the content and the views expressed.

## **TABLE OF CONTENTS**

EXECUTIVE	SUMMARY	V
SECTION 1	INTRODUCTION	1
SECTION 2	DATA AND METHODOLOGY	5
SECTION 3	DESCRIPTIVE ANALYSIS	9
SECTION 4	RESULTS	17
SECTION 5	SUMMARY AND CONCLUSION	31
REFERENC	ES	35
APPENDIX		37
LIST OF	TABLES	
Table 1	Overall incidence	9
Table 2a	Overall incidence by region	10
Table 2b	Regional variation	10
Table 3a	Overall incidence by sector	11
Table 3b	Sectoral variation of minimum wage workers	11
Table 4	Average characteristics of minimum wage workers by region in 2017 and 2018	12
Table 5	Characteristics of minimum wage workers by sector in 2017 and 2018	14
Table 6	Characteristics of treated and control by year	16
Table 7	Difference-in-Differences estimates of the effect of minimum wage change on hours worked	18
Table 8	Difference-in-Differences estimates of the effect of minimum wage change on hours worked by sector	19
Table 9	Difference-in-Differences estimates of the effect of minimum wage change on hours worked by region of work	21
Table 10	Difference-in-Differences estimates of the effect of minimum wage change on hours worked by region of residence	22
Table 11	Difference-in-Differences placebo estimates by sector	24
Table 12	Difference-in-Differences placebo estimates by region	25
Table 13	Difference-in-Differences hours estimates by sector	27

			1 ((		1 1	C + I	
IV	The	prevalence	and effect	on hours	worked o	t the minimum	wage in Ireland

Table 14	Difference-in-Differences hours estimates by region	28
Table A1	Robustness checks for sector	37
Table A2	Robustness checks for region	37
Table A3	Counties within NUTS 3 regions in 2018	38
Table A4	Region classification used in the report	38

#### **EXECUTIVE SUMMARY**

- This study examines the extent to which the profile of minimum wage (MW) employees, and the impacts of the 2018 increase in the national minimum wage rate, varies by sector and region in Ireland. The research highlights variations in the importance of MW employment across regions and sectors. It also highlights the potential importance of considering regional and sectoral issues in any process used to determine the appropriate MW rate.
- The incidence of MW employment remained constant, at 8.1 per cent of employees, in 2017 and 2018. This incidence varied substantially by region in 2018, from just 5 per cent of total employees in Dublin to approximately 10 per cent of employees in the Border, Midland, and Mid-West/South East regions.
- With respect to sector, approximately 30 per cent of all employees in the
  accommodation and food sector are minimum wage workers. In wholesale and retail,
  the incidence is approximately 16 per cent, with the figure for manufacturing
  standing at just over 5 per cent. The average incidence of MW employment across
  other sectors of the economy is 4.5 per cent.
- Minimum wage employees in Dublin tend to be younger, better educated, more likely
  to be employed on temporary contracts and have lower tenure than most other
  regions. While 22 per cent of Dublin MW workers are educated to tertiary level, the
  figure for the national average is 17 per cent. Minimum wage workers in Dublin are
  also more likely to be lone parents, non-Irish nationals and are less likely to be
  married relative to the average national minimum wage profile.
- Minimum wage employees in the manufacturing sector are more likely to be older, male, married, working full-time, on permanent contracts and have job tenure in excess of three years. These characteristics are suggestive of 'career minimum wage workers', as opposed to transient minimum wage employees, such as young, part-time workers who are also in education. This has important implications, as career minimum wage employees may be solely reliant on statutory MW increases to boost their earnings, whereas transient MW employees are likely to naturally progress up the wage distribution as they acquire further education, training and experience.
- Our analysis indicates that the 2018 MW increase did not lead to changes in the
  average hours worked of minimum wage employees generally. However, the
  research demonstrates that such aggregate analysis can mask heterogeneous
  impacts that can occur at the regional or sectoral level. In the Dublin and West
  regions, as well as in the manufacturing sector, the average hours worked of
  minimum wage employees fell in the six-month period following the minimum wage

increase. During the same period the average hours worked of non-minimum wage employees remained constant or increased slightly in these regions and sector.

- In the six-month period following the 2018 MW rise, average hours worked by MW employees, relative to non-MW employees, fell by 1.6 hours per week in Dublin and 1.7 hours per week in the West region. The sectoral models indicate that the average hours worked of minimum wage employees in the manufacturing sector, relative to non-MW employees, fell by 1.6 hours per week following the rate rise. The hours worked of minimum wage employees in the accommodation and food or wholesale and retail sectors were not affected.
- When the analysis was run for placebo periods, during which no change occurred in the MW rate, we found no impact on hours worked. This suggests that the observed impacts are likely related to the MW change, as opposed to diverging trends in the hours worked of minimum and non-minimum wage employees in particular sectors or regions prior to the rate change. However, we cannot fully discard this possibility due to the limitations in the number of placebo periods during which diverging trends can be tested.
- Our analysis indicates that the changes in hours worked mainly affected incumbent MW employees who were already employed in the period prior to the introduction of the 2018 rate rise. There is no evidence to support the view that the observed changes were due to compositional impacts whereby, for instance, more part-time employees entered (or left) employment in response to the rate rise, thereby altering the average hours worked of minimum wage employees in particular sectors or regions.
- The results suggest that the observed negative impacts on hours for Dublin and the West did not persist throughout 2018. While we detect an immediate reduction in hours worked of MW employees, relative to non-MW employees, in the first two quarters of 2018, the effects did not persist into the second half of the year.
- Our central results are robust to a range of specifications; however the data at hand
  do not allow us to identify the exact responses that brought about the observed
  impacts. There is a need for greater insights into exactly how firms and employees
  behave, following MW rate changes that generate a statistically significant change in
  hours worked, so that policy can be properly informed. This is particularly the case in
  instances such as those in the manufacturing sector, where impacts are observed
  among individuals who appear to be career minimum wage employees.
- The finding of possible heterogeneous impacts arising from the 2018 minimum wage change suggests that the existence of such variations should be monitored, particularly given that any impacts could become more persistent during periods of lower economic growth.

#### Introduction

A Low Pay Commission was established in Ireland in 2015 to advise the Irish government on the appropriate level of the national minimum wage. In doing so, the Low Pay Commission seeks a minimum wage that is fair and sustainable and helps as many low paid workers as possible, without creating adverse employment effects. There has been much debate, and often-conflicting evidence, about the impact of minimum wage changes on employment outcomes. Some recent international studies find adverse employment effects (e.g. Sabia et al., 2016; Galán and Puente, 2015; Meer and West, 2016), while others find little to no negative effect (e.g. Cengiz et al., 2018; Hoffman, 2016; Schmitt, 2015; Belman et al., 2015). A related strand of literature looks at the distributional impacts of minimum wage changes. While minimum wage changes have been found to reduce hourly wage inequality, the effect on household income is limited, as a large number of minimum wage workers are located in the top half of the household income distribution (Redmond et al., 2018; Logue and Callan, 2016; MaCurdy, 2015). However, for the US, Dube (forthcoming) shows that higher minimum wages are an effective means of increasing incomes of families in the bottom half of the wage distribution and thereby reducing the poverty rate.

Recent work has highlighted the importance of accounting for possible heterogeneous effects among different subgroups of workers in the population, which could be overlooked when studying overall employment outcomes. For example, Dickens et al. (2015) find no minimum wage related employment effects for full-time workers in the UK, however a negative impact is found when focusing on part-time females only. Likewise, minimum wage effects have been found to vary by age (Liu et al., 2016; Galán and Puente, 2015) and firm-type, i.e. production or non-production (Del Carpio et al., 2015).

It is also important to explore heterogeneity in the type of minimum wage employee. While minimum wages are typically set at a national level, such an aggregate approach potentially risks exposing groups of employees located in particular sectors or regions to increased risk of negative impacts. Certain groups of employees may find themselves in persistent low wage employment and are therefore solely reliant on statutory minimum wage changes to boost their incomes. We can refer to such workers as 'career minimum wage employees'. For others, such as students working part-time while in third-level education, minimum wage employment is likely to be short term in nature. Such employees will naturally transition to higher pay as they acquire skills, education and experience, and are therefore not reliant on minimum wage increases. Redmond

temporary contracts.

et al. (2018) study the transition patterns of minimum wage workers in Ireland and find that, while the minimum wage typically acts as a stepping stone to higher pay, a substantial minority of minimum wage workers remain on the minimum wage for longer periods of time. Irish nationals, older workers, those with higher levels of education, full-time employees and those on permanent contracts are more likely to transition to higher pay compared to non-nationals, younger persons, those with lower educational attainment, part-time workers and those on

The heterogeneity of minimum wage workers, and the associated effects, underpins the importance of focusing not just on the full population of minimum wage workers, but on separate subgroups. In this paper, we study both regional and sectoral variation of minimum wage employment using a unique feature of the Irish data that comes from a question that was added to Ireland's Labour Force Survey in Quarter 2 of 2016, which directly asks employees whether they are on the minimum wage. Minimum wage employment is found to be heavily concentrated in the accommodation and food and wholesale and retail sectors. These two sectors alone account for 55 per cent of minimum wage workers. The characteristics of minimum wage employees vary by sector. Minimum wage workers in the manufacturing sector are shown to have characteristics consistent with 'career minimum wage employees'. Compared to the average profile of minimum wage employees in Ireland, those employed in manufacturing tend to be older, married, predominantly male and work full-time hours. Moreover, they are heavily concentrated outside of Dublin, in the Midland and Border/Mid-East regions.

Following recommendations from the Low Pay Commission, the Irish NMW increased in January 2018 from €9.25 per hour to €9.55 per hour. When estimating the impact of the 2018 minimum wage increase on the hours worked of all minimum wage employees in Ireland, we find no effects. However, our sectoral and regional analysis uncovers significant effects, which underpin the importance of accounting for heterogeneous impacts of minimum wage policies. With regard to region, we detect statistically significant hours effects for minimum wage workers in Dublin, with a reduction, relative to non-MW employees, of 1.6 hours per week. There is also evidence to indicate a reduction of 1.7 hours per week for MW employees, relative to non-MW employees, in the West region. Minimum wage employees in the other four regions experience no significant changes to their hours. With regard to sector, we detect a strong negative impact for manufacturing employees, who experience a decline of 1.6 hours per week.

Our findings highlight the importance of unpacking minimum wage statistics to uncover regional and sectoral effects that may be hidden within broad, population-level studies. Despite the large body of research studying minimum wages, regional

studies receive relatively little attention in the minimum wage literature. The few studies that do exist further support the importance of this type of research. Wang et al. (2019), using a US restaurant industry panel from 1990 to 2006, find substantial heterogeneity across groups and regions including positive and negative impacts. Williams (1993) finds substantial regional variation in the effect of minimum wages on employment in the US, which is largely unobserved when focusing on national-level data. Gilbert et al. (2001) find differences in the characteristics of rural and urban minimum wage workers in the UK, as well as differences in the potential impacts on earnings inequality.

The remainder of the paper proceeds as follows. In Section 2, we describe the data and outline our methodology. Section 3 presents some relevant descriptive statistics of minimum wages employees across regions and sectors. Section 4 presents our main results and Section 5 concludes.

### Data and methodology

Our data come from the Irish Labour Force Survey (LFS), which is a large-scale, quarterly survey of households in Ireland. The LFS is carried out by the Central Statistics Office (CSO) and provides the official measures of unemployment in Ireland, as well as a variety of other quarterly labour force statistics for the working age population of people aged 15 and over. The LFS is a rich dataset on individual characteristics including age, sex, region of residence, nationality, education and a range of other variables related to the individual's labour market status. Of particular relevance to our study is a question, unique to the Irish Labour Force Survey, which asks individuals whether or not they are on the minimum wage. The question was added to the LFS in Quarter 2 of 2016. It asks individuals whether they earn (a) less than the minimum wage, (b) exactly the minimum wage or (c) more than the minimum wage. The analysis in this paper focuses on the working population of people aged 15 and over.<sup>2</sup>

As we are examining the effect of the 2018 minimum wage increase on hours worked, we focus on the years 2017 and 2018. We use usual hours worked in the analysis but our results are robust to using actual hours worked. The NUTS3 regional classification is used for a person's region of employment. This classification divides Ireland into the following eight regions: Border, West, Mid-West, South East, South West, Dublin, Mid-East and the Midland. In Quarter 1, 2018, the NUTS3 boundaries were amended slightly such that county Louth changed from being included in the Border region to being included in the Mid-East region and South Tipperary changed from the South East region to the Mid-West region. To avoid these changes impacting our results, we group the Border and Mid-East regions together in to one region and we also group the South East and Mid-West regions together. Thus, we focus our analysis on the following six regions: West, South East/Mid-West, Dublin, Border/Mid-East, South West, and the Midland. We study three sectors which represent the highest concentration of minimum wage employment – wholesale and retail, accommodation and food, and manufacturing. More than half of all minimum wage employees are located in the first two sectors, with 8 per cent located in the manufacturing sector. The remaining minimum wage employees are located across a very broad range of sectors. As the numbers within each specific sector are too low to carry out any

Prior to the introduction of this question, ascertaining who was and was not a minimum wage employee was difficult due to the lack of precise wage data. To overcome this, researchers used a combination of income decile and hours worked to impute minimum wage status (see McGuinness and Redmond, 2019).

<sup>&</sup>lt;sup>2</sup> The question was asked to all employees.

meaningful analysis, we group these remaining minimum wage employees into an 'other' category.

When evaluating the effect of the 2018 minimum wage increase, it is not sufficient to compare the hours worked of minimum wage employees before and after the increase, as hours may have changed even in the absence of a policy change. To overcome this, we use a difference-in-differences (DiD) strategy that compares the change in hours worked among minimum wage employees (our treatment group) to the change in hours worked of a control group consisting of non-minimum wage employees. Therefore, the DiD strategy subtracts the change in hours, pre- and post-policy change, among the control group, from the change in hours worked among the treatment group over the same period. If the change in hours for minimum wage employees (the treatment group) is larger than the change in hours for non-minimum wage employees (the control group), then it is likely that the hours effect is attributable to the policy change. In order to avoid seasonal effects, we compare Quarters 1 and 2 in 2017 with Quarters 1 and 2 in 2018. More formally, the DiD estimator can be implemented with the following regression,

$$Y_{i,t,w} = \beta_{1w} + \beta_{2w} Y ear_t + \beta_{3w} T_{i,w} + \beta_{4w} Y ear_t * T_{i,w} + X_{i,t,w}' \beta_{5w} + \varepsilon_{i,t,w}$$
(1)

In Equation (1),  $Y_{i,t,w}$  represents hours worked for individual i in year t and region w. Year<sub>t</sub> is a dummy variable which equals 1 for observations in Quarters 1 and 2 in 2018, i.e. the post-policy change period, and equals 0 for observations in Quarters 1 and 2 in 2017, pre-policy change.  $T_{i,w}$  is a treatment dummy variable which equals 1 if the individual is a minimum wage employee and zero if a higher paid employee. The interaction term  $Year_t *T_{i,w}$  represents the estimated treatment effect. We also include a vector of additional controls,  $X_{i,t,w}$  which include age, education, gender, quarter, and a binary variable denoting whether the individual has children. The additional variables help to control for any compositional changes in the group of minimum wage employees over the two time periods, as well as improving the precision of the estimates.

While the LFS data do not contain precise wage information (except for minimum wage employees), workers are allocated to income deciles. However, this variable is poorly populated and information is missing for the majority of workers. Nevertheless, for the subset of employees with decile information, we observe that minimum wage employees are not represented in the higher deciles, namely deciles 9 and 10. This is to be expected, given the relatively low hourly wage rate of minimum wage employees. Therefore, to improve comparability between the

When we examine the impact of the minimum wage change across sectors, w denotes the sector.

treatment and control groups, we remove from the control group any workers in deciles 9 or 10, for cases where decile information is available.

The existence of the minimum wage question in the LFS allows us a unique opportunity to accurately measure the incidence of the MW and effectively evaluate the impact of rate changes on outcome variables, such as hours worked, using empirical techniques that are widely used in the international literature. Nevertheless, as in all counterfactual studies of this nature, some caution is required in interpreting the results. While the use of a control group in the DiD estimator allows us to account for factors, other than the NMW change, which may have impacted the hours of all employees, researchers can never fully guarantee that other factors were not in some way impacting the results. However, for this to be the case, such factors would have to affect the treatment and control groups in different ways. When interpreting the results, it is also worth noting that we cannot easily identify the behavioural changes among either employers or employees that lead to an impact on hours. Nevertheless, we have undertaken a series of robustness tests to verify the validity of our findings. Our robustness tests also provide information on the mechanisms behind hours changes, by disentangling the hours impacts on existing MW employees from any hours changes driven by new MW employees entering (or MW workers leaving) employment after the 2018 rate rise.

### **Descriptive analysis**

Table 1 shows the incidence of national minimum wage employment in Ireland. In both 2017 and 2018, 8.1 per cent of employees were minimum wage employees. When looking at the incidence of minimum wage employment, we define minimum wage employees as those earning on or below the national minimum wage. During the years covered in this study, 2017 and 2018, there existed subminimum wage rates for the following workers; young workers under 18 years of age, individuals who were in structured training during working hours, and inexperienced workers in their first two years of employment.<sup>4</sup> The sub-minimum wage rates range from between 70 to 90 per cent of the full rate. However, subminimum wage employment is very rare, making up approximately one percentage point of the overall incidence (of 8 per cent). Therefore, due to small sample sizes, carrying out separate analysis on the sub-minimum wage group is not feasible. However, it does make sense to include both sub-minimum and minimum wage employees in the same group, as they are all minimum wage employees. Furthermore, an increase in the statutory minimum wage affects both types of workers, as sub-minimum and minimum wage rates both increase simultaneously.

TABLE 1 OVERALL INCIDENCE

	2017	2018	Overall
MW Workers	0.081	0.081	0.081
Observations	50,302	47,436	97,738

Source: Irish Labour Force Survey 2017 and 2018.

Notes: LFS for all quarters in 2017 and 2018. Minimum Wage workers calculated as a percentage of all employees.

Table 2a indicates that the incidence of minimum wage employment varies substantially by region. In 2018, just 5 per cent of employees in Dublin were earning the minimum wage, compared to approximately 10 per cent in the Mid-West/South East, Border/Mid-East and Midland regions. The incidence in the South West and West regions was approximately 9 per cent. The regional statistics relate to the area where the individual works, rather than where they live. However, just 8 per cent of the sample was found to work outside the region where they live. Table 2b shows the regional distribution of minimum wage employees. Of the full sample of minimum wage employees nationally, there is a large concentration located in Dublin (23 per cent), the Mid-West/South East (23 per cent) and the Border/Mid-East (21 per cent) regions. The South West, West and

<sup>&</sup>lt;sup>4</sup> New criteria for sub-minimum rates, based solely on age, were implemented in March 2019.

Midland regions account for 17 per cent, 10 per cent and 6 per cent, respectively, of all minimum wage employees in Ireland.

TABLE 2A OVERALL INCIDENCE BY REGION

	Border/ Mid-East	West	Mid-West/ South East	South West	Dublin	Midland
2017 MW Workers	0.100	0.086	0.107	0.095	0.053	0.099
ZU17 IVIVV VVOIKEIS	(n=7,837)	(n=4,620)	(n=8,866)	(n=7,250)	(n=18,155)	(n=2,231)
2018 MW Workers	0.105 (n=8,163)	0.091 (n=4,124)	0.107 (n=9,977)	0.085 (n=7,330)	0.050 (n=16,999)	0.102 (n=2,464)
Overall MW Workers (2017 and 2018)	0.103	0.088	0.107	0.090	0.052	0.100
Observations	16,000	8,744	16,843	14,580	35,154	4,695

Source: Irish Labour Force Survey 2017 and 2018.

Notes: LFS for all quarters in 2017 and 2018. Minimum Wage workers calculated as a percentage of all employees. Region of work is

used in the analysis.

TABLE 2B REGIONAL VARIATION

	Border/ Mid-East	West	Mid-West/ South East	South West	Dublin	Midland
2017 MW Workers	0.195 (n=782)	0.099 (n=396)	0.236 (n=946)	0.173 (n=692)	0.242 (n=971)	0.055 (n=220)
2018 MW Workers	0.226 (n=860)	0.099 (n=376)	0.224 (n=855)	0.164 (n=624)	0.222 (n=847)	0.066 (n=251)
Overall MW Workers (2017 and 2018)	0.21	0.099	0.230	0.168	0.232	0.060
Observations	1,642	772	1,801	1,316	1,818	471

Source: Irish Labour Force Survey 2017 and 2018.

Notes: LFS for all quarters in 2017 and 2018. Minimum Wage workers calculated as a percentage of all employees. Region of work is used in the analysis.

We also observe substantial variation in the incidence of minimum wage employment across sectors. Table 3a shows that, in 2018, 16 per cent of all employees in the wholesale and retail sector and 30 per cent of all employees in the accommodation and food sector were minimum wage employees. Table 3b indicates that these two sectors alone account for 54 per cent of all minimum wage employees in Ireland. Approximately 5 per cent of all manufacturing workers are minimum wage employees, accounting for 8 per cent of all minimum wage employment in Ireland. The remaining 37 per cent of minimum wage workers in Ireland are widely distributed across other sectors of the labour market.<sup>5</sup>

The other sectors in which minimum wage employees made up more than 1 per cent of all employees included the following: Crop and animal production, hunting and related service activities; Construction of buildings; Specialised construction activities; Land transport and transport via pipelines; Services to buildings and landscape activities; Public administration and defence (compulsory social security); Education; Human health activities; Residential care activities; Social work activities without accommodation; Sports activities and amusement and recreation activities; and Other personal service activities.

	Wholesale and Retail	Accommodation and Food	Manufacturing	Other
2017 MW Workers	0.160 (n=7,146)	0.292 (n=3,694)	0.053 (n=6,146)	0.045 (n=33,225)
2018 MW Workers	0.160 (n=6,577)	0.304 (n=3,405)	0.054 (n=5,675)	0.046 (n=31,657)
Overall MW Workers (2017 and 2018)	0.16	0.298	0.053	0.045
Observations	13,723	7,099	11,821	64,882

Source: Irish Labour Force Survey 2017 and 2018.

Notes: LFS for all quarters in 2017 and 2018. Minimum Wage workers calculated as a percentage of all employees.

TABLE 3B SECTORAL VARIATION OF MINIMUM WAGE WORKERS

	Wholesale and Retail	Accommodation and Food	Manufacturing	Other
2017 MW Workers	0.282 (n=1,141)	0.266 (n=1,078)	0.080 (n=324)	0.371 (n=1,503)
2018 MW Workers	0.274 (n=1,052)	0.269 (n=1,035)	0.080 (n=308)	0.376 (n=1,446)
Overall MW Workers (2017 and 2018)	27.8	26.8	0.080	37.4
Observations	2,193	2,113	632	2,949

Source: Irish Labour Force Survey 2017 and 2018.

Notes: LFS for all quarters in 2017 and 2018. Minimum Wage workers calculated as a percentage of all employees.

Table 4 examines the mean characteristics of MW employees by region across a range of dimensions including gender, age, education, marital status, household composition, hours worked, contractual status, firm size and sector. The final column of the table provides the basis for comparison by listing the average characteristics of the full national sample of minimum wage employees. We undertake t-tests that indicate whether the regional averages vary in a statistically significant way from the national average. It is apparent that minimum wage employees' characteristics in Dublin are quite different to the characteristics of minimum wage employees at the national average level. Firstly, those on the minimum wage in Dublin are younger and better educated than in the rest of the country. The average age of minimum wage employees in Dublin is 29, compared to approximately 31 elsewhere. While 22 per cent of the Dublin workers are educated to tertiary level, the figure for the national average is 17. Minimum wage workers in Dublin are also more likely to be lone parents, non-Irish nationals, have temporary contracts, have lower tenure, and are less likely to be married relative to the average national minimum wage profile. It is also notable that minimum wage workers in the manufacturing sector are predominantly located outside of Dublin; just 3 per cent of minimum wage employees in Dublin are in the manufacturing sector, whereas the corresponding figure for other regions is between 9 and 10 per cent. Other regional based differences include: (i) higher proportions of married MW employees in the West and Midland regions, (ii) lower proportions of MW employees working part-time in the West region, and (iii) a lower percentage of MW employees employed in small businesses in Dublin and the South West.

TABLE 4 AVERAGE CHARACTERISTICS OF MINIMUM WAGE WORKERS BY REGION IN 2017 AND 2018

	Border/Mid- East	West	Mid- West/South East	South West	Dublin	Midland	Ireland
Age	31.54	32.70***	31.60	30.80	28.76***	33.10***	31.01
Female	0.55	0.56	0.58**	0.56	0.51***	0.56	0.55
Lone parent	0.22	0.17**	0.19*	0.18*	0.26***	0.20	0.21
Single household	0.12***	0.24***	0.16	0.14	0.17	0.14	0.16
Married	0.25	0.31***	0.23	0.24	0.20***	0.30***	0.24
Irish	0.84	0.79***	0.87***	0.84	0.78***	0.81	0.83
Nationality other	0.02***	0.04	0.02***	0.03***	0.10***	0.05	0.04
EU	0.14	0.17	0.11	0.13	0.12	0.14	0.13
No children	0.63	0.65	0.63	0.63	0.63	0.62	0.63
1 to 2 children	0.32	0.30	0.33	0.32	0.33	0.31	0.32
More than 2 children	0.05	0.05	0.05	0.04	0.04	0.07*	0.05
Low ISCED	0.25**	0.24	0.23	0.23	0.18***	0.22	0.22
Medium ISCED	0.62	0.60	0.62	0.59	0.60	0.63	0.61
High ISCED	0.13***	0.17	0.14***	0.18	0.22***	0.15	0.17
Temporary Contract	0.32	0.27***	0.32	0.34	0.36**	0.32	0.33
Hours Worked	23.59*	25.02	23.78	24.93	24.74	23.89	24.34
Part time	0.62***	0.54***	0.59	0.58	0.58	0.61	0.59
Firms with ≤10 workers	0.35	0.37	0.42***	0.33**	0.30***	0.44***	0.36
Approx 1 year tenure	0.52	0.47***	0.51**	0.55	0.63***	0.49**	0.54
Approx 2 years tenure	0.16**	0.14	0.13	0.14	0.13	0.14	0.14
Approx 3 years tenure	0.08	0.10**	0.09*	0.07	0.07	0.08	0.08
More than 3 years tenure	0.24	0.29***	0.27**	0.24	0.17***	0.30***	0.24
Wholesale and Retail	0.28	0.27	0.26	0.30*	0.28	0.25	0.28
Accommodation and Food	0.26	0.28	0.27	0.26	0.27	0.28	0.27
Manufacturing	0.10***	0.09	0.09*	0.09	0.03***	0.10*	0.08
Other	0.36	0.36	0.38	0.35*	0.41***	0.36	0.37
Observations	1,642	772	1,801	1,316	1,818	471	7,901

Source: Irish Labour Force Survey 2017 and 2018.

Note:

T-test of differences in sample means between each region and Ireland performed with \*\*\* denoting significance at 0.01 level. \*\* at 0.05 and \* at 0.10.

Table 5 examines the mean characteristics of minimum wage employees by sector of employment. Minimum wage employees employed in the manufacturing sector look very different from minimum wage workers employed in other sectors of the economy across a range of characteristics. In particular, compared to the national average, minimum wage employees in manufacturing are more likely to be older, male, non-Irish nationals, married, working full-time hours, in permanent positions and in firms with more than ten workers. For example, 79 per cent of workers in the manufacturing sector are on permanent contracts and 79 per cent are in fulltime positions, which compares to national averages for MW employees of 67 and 41 per cent respectively. The characteristics of minimum wage employees in the manufacturing sector are therefore consistent with 'career minimum wage employees'. This has potentially important policy implications. Career minimum wage employees will be reliant on statutory increases in the national minimum wage to increase their incomes. It is also likely that their income from work represents a large component of their household income, meaning that any increase in the minimum wage could improve the standard of living of these households. On the other hand, the minimum wage is likely to be a short-term situation for young, part-time workers who are highly educated. These transient minimum wage workers will not be as reliant on statutory increases in the minimum wage to boost their incomes as they may naturally progress up the wage scale as they acquire further education, training and experience.

Other sectoral variations of note include (i) a higher than average share of Irish MW employees in wholesale and retail, (ii) above average shares of part-time MW employees in both accommodation and food and wholesale and retail, (iii) lower employment tenure in accommodation and food and (iv) a higher incidence of female MW workers in the accommodation and food and wholesale and retail sector.

TABLE 5 CHARACTERISTICS OF MINIMUM WAGE WORKERS BY SECTOR IN 2017 AND 2018

	Wholesale and Retail	Accommodation and Food	Manufacturing	Other	All Sectors
	Mean	Mean	Mean	Mean	Mean
Age	28.28***	27.42***	34.05***	34.92***	31.01
Female	0.58***	0.62***	0.34***	0.52***	0.55
Lone parent	0.20	0.24***	0.15***	0.21	0.21
Single household	0.12***	0.17**	0.18*	0.16	0.16
Married	0.19***	0.19***	0.33***	0.29***	0.24
Irish	0.89***	0.76***	0.69***	0.85***	0.83
Nationality other	0.02***	0.07***	0.06*	0.04*	0.04
EU	0.09***	0.16***	0.25***	0.11**	0.13
No children	0.62	0.64	0.63	0.64	0.63
1 to 2 children	0.33	0.32	0.33	0.31	0.32
More than 2 children	0.05	0.04	0.04	0.05	0.05
Low ISCED	0.21	0.21	0.17***	0.25***	0.22
Medium ISCED	0.65***	0.63	0.68***	0.56***	0.61
High ISCED	0.14***	0.17	0.15	0.19***	0.17
Temporary Contract	0.33	0.36**	0.21***	0.33	0.33
Hours Worked	21.64***	20.34***	34.48***	27.03***	24.34
Part-time	0.68***	0.71***	0.21***	0.51***	0.59
Firms with ≤ 10 workers	0.37	0.28***	0.16***	0.46***	0.36
Approx 1 year tenure	0.55	0.61***	0.52	0.49***	0.54
Approx 2 years tenure	0.15	0.14	0.11**	0.14	0.14
Approx 3 years tenure	0.07	0.08	0.09	0.08	0.08
More than 3 years tenure	0.23	0.17***	0.29***	0.29***	0.24
Border/Mid-East	0.21	0.20	0.27***	0.20	0.21
West	0.10	0.10	0.11	0.09	0.10
Mid-West/South East	0.22	0.23	0.26*	0.23	0.23
South West	0.18	0.17	0.18	0.16	0.17
Dublin	0.24	0.24	0.10***	0.25**	0.23
Midland	0.06	0.06	0.08*	0.06	0.06
Observations	2,193	2,113	632	2,949	7,901

Source: Irish Labour Force Survey 2017 and 2018.

Note: T-test of differences in sample means between each sector and overall sectors performed with \*\*\* denoting significance at 0.01 level, \*\* at 0.05 and \* at 0.10.

Finally, Table 6 compares the characteristics of minimum wage employees (the treatment group) and those earning above the minimum wage (the control group). Not surprisingly, and consistent with other research by Redmond et al. (2018), some differences emerge in the average characteristics of the two groups. Relative

to the control group, minimum wage employees are much more likely to be younger, less well educated, lone parents, on temporary contracts, non-Irish, in part-time employment, located in smaller firms and have shorter employment tenures. As shown in Table 6, while there are some significant differences in the characteristics of the treatment and control groups, these differences are stable over time. This is important for our difference-in-differences methodology, which subtracts the change in hours from 2017 to 2018 in the control group, from the change in hours over the same period in the treated group. If there were substantial changes over time in the average characteristics of one group only, such that the differences in the characteristics of both groups were not stable over time, this would raise the likelihood that any observed hours effect was attributable to changes in characteristics as opposed to the MW change. However, Table 6 demonstrates that this does not appear to be an issue in our analysis.

TABLE 6 CHARACTERISTICS OF TREATED AND CONTROL BY YEAR

	Control 2017	Treated 2017	Control 2018	Treated 2018
	Mean	Mean	Mean	Mean
Age	41.29***	31.46	42.07***	31.39
Female	0.52***	0.58	0.52**	0.55
Lone parent	0.11**	0.21	0.11***	0.21
Single household	0.15***	0.18	0.14**	0.16
Married	0.56***	0.26	0.59***	0.24
Irish	0.88***	0.77	0.89***	0.85
Nationality other	0.03***	0.05	0.02***	0.03
EU	0.10**	0.17	0.09***	0.12
No children	0.57**	0.59	0.57***	0.67
1 to 2 children	0.35	0.35	0.34***	0.29
More than 2 children	0.08***	0.06	0.09***	0.04
Low ISCED	0.11***	0.21	0.11***	0.21
Medium ISCED	0.37***	0.60	0.37***	0.62
High ISCED	0.52***	0.19	0.51***	0.17
Temporary Contract	0.06***	0.27	0.07***	0.32
Hours Worked	34.03***	23.45	34.73***	24.41
Part-time	0.19***	0.59	0.19***	0.59
Firms with ≤ 10 workers	0.23***	0.38	0.20***	0.35
Approx 1 year tenure	0.18***	0.49	0.19***	0.48
Approx 2 years tenure	0.09***	0.17	0.09***	0.15
Approx 3 years tenure	0.07	0.08	0.08***	0.10
More than 3 years tenure	0.66***	0.25	0.64***	0.26
Wholesale and Retail	0.13***	0.28	0.13***	0.27
Accom. and Food	0.06***	0.27	0.06***	0.27
Manufacturing	0.12***	0.08	0.12***	0.08
Other	0.68***	0.37	0.69***	0.38
Observations	24,623	2,101	19,418	2,031

Source: Irish Labour Force Survey 2017 and 2018.

Notes: This analysis includes Quarters 1 and 2 only. Top two income deciles for control group are also excluded. T-test of differences in sample means between treated and control performed with \*\*\* denoting significance at 0.01 level, \*\* at 0.05 and \* at 0.10.

#### **SECTION 4**

#### **Results**

We next present the results from our empirical methodology. It is important to stress that it is almost impossible to be certain that we are identifying a causal impact of a change in the 2018 MW rate on hours worked. Nevertheless, we adopt an empirical strategy that attempts to remove particular forms of bias that can potentially impact our estimates, thus allowing us to point towards potentially causal effects with a greater degree of certainty.

In Table 7 we estimate Equation (1) for all employees, and then separately for temporary employees.<sup>6</sup> This is done on the grounds that McGuinness and Redmond (2018) found a statistically significant impact following the introduction of the 2016 MW rate rise that was more pronounced among temporary employees.<sup>7</sup> We estimate each model both with and without control variables to test the sensitivity of the estimate to possible compositional changes among the characteristics of the treatment and control groups. The DiD coefficients are not statistically significant in all four models. This indicates that the 2018 increase in the national minimum wage did not lead to a change in the overall average hours worked of minimum wage employees in Ireland.<sup>8</sup>

The sample sizes will be smaller than for the descriptive analysis, as the models are run using two quarters of data from each year.

We can estimate the full model on the subset of temporary workers. However, due to restricted sample sizes, we cannot separately estimate temporary contract models for each region and sector.

The coefficients associated with the additional control variables behave as expected, indicating that the average number of hours worked is lower for females and those with children. Those with higher levels of education work more hours.

TABLE 7 DIFFERENCE-IN-DIFFERENCES ESTIMATES OF THE EFFECT OF MINIMUM WAGE CHANGE ON HOURS WORKED

	(1)	(2)	(3)	(4)
VARIABLES	Overall	Overall	Temporary workers	Temporary workers
DiD	0.165	-0.0652	0.996	1.189
	(0.341)	(0.323)	(0.893)	(0.858)
Time	0.113	0.190**	-0.101	-0.171
	(0.0983)	(0.0927)	(0.488)	(0.468)
Minimum Wage	-9.726***	-8.852***	-8.593***	-6.699***
	(0.239)	(0.234)	(0.646)	(0.650)
Age		-0.0325***		-0.00649
		(0.00381)		(0.0144)
Female		-6.911***		-5.244***
		(0.0892)		(0.400)
Children (binary)		-0.560***		-0.0259
		(0.0899)		(0.421)
Medium ISCED (ref = low ISCED)		2.005***		4.611***
		(0.150)		(0.593)
High ISCED		4.454***		9.129***
		(0.150)		(0.622)
2 <sup>nd</sup> Quarter		0.0751		0.989**
		(0.0883)		(0.392)
Constant	35.29***	37.37***	28.34***	24.59***
	(0.0655)	(0.241)	(0.346)	(0.902)
Observations	46,460	45,520	3,614	3,585
R-squared	0.064	0.189	0.083	0.168

Source:

Irish Labour Force Survey 2017 and 2018.

Notes:

The DiD model compares Q1/Q2 2017 with Q1/Q2 2018. Standard errors in parentheses with significance levels \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

Tables 8 and 9 present the results for Equation (1), estimated separately by region and sector. While no overall impact was observed for all MW employees, there is evidence of heterogeneous impacts across sectors and regions. The sectoral models in Table 8 indicate that only minimum wage employees in the manufacturing sector recorded a fall in the average hours worked. Relative to non-minimum wage employees in the same sector, minimum wage employees in manufacturing experienced a reduction of 1.6 hours per week. No significant effects were detected for minimum wage employees in the accommodation and food and wholesale and retail sectors. The hours of minimum wage employees in the 'other' sectors increased by 0.9 hours per week, relative to non-minimum wage employees in the same sectors, following the 2018 rate rise. However, the impact is only marginally statistically significant. The finding with regard to manufacturing is potentially concerning given the evidence presented earlier which indicates that

such employees are likely to be career minimum wage employees who are working in full-time, permanent positions. A drop in hours for these types of employees may be more damaging than a drop in hours for a young, highly educated, part-time employee, for which minimum wage employment may simply be a temporary stepping stone to higher pay.

TABLE 8 DIFFERENCE-IN-DIFFERENCES ESTIMATES OF THE EFFECT OF MINIMUM WAGE CHANGE ON HOURS WORKED BY SECTOR

	(1)	(2)	(3)	(4)
VARIABLES	Wholesale and Retail	Accommodation and Food	Manufacturing	Other
DiD	-0.432	-0.0685	-1.621**	0.938*
	(0.668)	(0.827)	(0.775)	(0.509)
Time	-0.592**	-0.622	0.718***	0.328***
	(0.268)	(0.445)	(0.180)	(0.111)
Minimum Wage	-9.109***	-8.196***	-2.733***	-7.560***
	(0.481)	(0.597)	(0.571)	(0.366)
Age	0.0827***	0.128***	-0.0184**	-0.0973***
	(0.0103)	(0.0158)	(0.00821)	(0.00472)
Female	-6.760***	-3.882***	-3.992***	-7.212***
	(0.246)	(0.379)	(0.187)	(0.110)
Children (binary)	-0.00495	-0.505	-0.258	-0.945***
	(0.248)	(0.383)	(0.176)	(0.110)
Medium ISCED (ref = low ISCED)	3.737***	3.828***	0.638**	1.709***
	(0.365)	(0.547)	(0.288)	(0.196)
High ISCED	6.712***	8.176***	2.106***	3.753***
	(0.400)	(0.595)	(0.290)	(0.192)
2 <sup>nd</sup> Quarter	-0.0941	0.230	0.179	0.0867
	(0.244)	(0.373)	(0.174)	(0.108)
Constant	30.41***	25.93***	39.30***	40.79***
	(0.621)	(0.925)	(0.494)	(0.307)
Observations	6,563	3,304	5,480	30,105
R-squared	0.250	0.236	0.104	0.168

Source: Irish Labour Force Survey 2017 and 2018.

Notes: The DiD model compares Q1/Q2 2017 with Q1/Q2 2018. Standard errors in parentheses with significance levels \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

With respect to region, following the 2018 MW increase, the average weekly hours worked by minimum wage employees, relative to non-minimum wage employees in the same region, fell by 1.6 hours in Dublin and 1.7 hours in the West region. However, the results for the West region are only marginally statistically significant. There were no statistically significant effects detected in any of the

other four regions. These results are robust to the inclusion of additional controls for sector in the regional models, and for region in the sectoral models.

The DiD estimates compare changes in the treatment group to changes in the control group. As such, hours changes in either the treatment or control group, or some combination of both, can drive the results. For example, the negative hours effect observed for MW employees could be due to hours worked of MW employees declining, while hours worked of non-MW employees remained the same. On the other hand, we could get a negative hours effect if the hours worked of MW employees stayed constant, but the hours of non-MW workers increased. To disentangle what is driving our results, we look at how the average hours of both groups actually changed over the two periods. For the West region, the effect is fully driven by a reduction in hours of MW employees: over the two time periods, the hours worked of MW employees fell by approximately 1.7 hours per week (or 6 per cent), while the hours of non-MW employees remained stable. For Dublin, the effect was also primarily driven by a fall in hours of MW employees. Over the two time periods, the hours worked of MW employees in Dublin fell by 3 per cent, with a simultaneous increase of just under 1 per cent for hours of non-MW workers. For manufacturing, there was a 3.5 per cent decline in the hours worked of MW employees, while non-MW employees in the manufacturing sector saw an increase in hours of approximately 1.5 per cent. What these results show is that the negative hours effects detected in our DiD analysis were either fully, or predominantly, driven by actual declines in the average hours worked of MW employees.

Our analysis is based on the individual's region of work. However, in order to assess the extent to which the choice of regional variable (i.e. region of work or region of residence) influences our results, we re-estimate our models using region of residence. The results are shown in Table 10. Using region of residence, we again see a fall in the hours worked of minimum wage employees in Dublin and the West. The results in this specification for Dublin and the West are of a slightly higher magnitude and of increased statistical significance in comparison to the region of work specification. There is also a weakly statistically significant fall in hours in the Midland region, as well as marginally significant increases in hours in the Mid-West/South East and South West regions. Nevertheless, this robustness test tends to confirm the decline in hours worked in both Dublin and the West regions among minimum wage workers following the introduction of the 2018 rate.

TABLE 9 DIFFERENCE-IN-DIFFERENCES ESTIMATES OF THE EFFECT OF MINIMUM WAGE CHANGE ON HOURS WORKED BY REGION OF WORK

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	Border/ Mid-East	West	Mid-West/ South East	South West	Dublin	Midland
DiD	-0.195	-1.693*	0.730	1.260	-1.621**	-1.332
	(0.748)	(1.010)	(0.698)	(0.791)	(0.636)	(1.406)
Time	0.438*	0.216	0.158	0.361	0.303**	0.693
	(0.237)	(0.293)	(0.231)	(0.242)	(0.149)	(0.429)
Minimum Wage	-9.013***	-7.041***	-8.720***	-8.699***	-8.903***	-8.914***
	(0.569)	(0.758)	(0.511)	(0.570)	(0.432)	(1.023)
Age	-0.0294***	-0.0788***	-0.0263***	-0.0403***	-0.0205***	-0.0189
	(0.00964)	(0.0122)	(0.00941)	(0.00997)	(0.00624)	(0.0180)
Female	-7.220***	-5.623***	-7.667***	-8.037***	-5.873***	-6.372***
	(0.230)	(0.285)	(0.221)	(0.233)	(0.144)	(0.418)
Children (binary)	-0.539**	-0.583**	-0.228	-0.869***	-0.647***	0.517
	(0.232)	(0.284)	(0.222)	(0.236)	(0.146)	(0.419)
Medium ISCED (ref = low ISCED)	2.619***	1.308***	2.299***	2.204***	1.963***	0.929
	(0.346)	(0.498)	(0.346)	(0.399)	(0.265)	(0.670)
High ISCED	4.160***	3.293***	3.800***	4.524***	5.267***	2.538***
	(0.355)	(0.506)	(0.355)	(0.397)	(0.256)	(0.705)
2 <sup>nd</sup> Quarter	-0.157	0.441	0.138	0.0254	0.0227	0.608
	(0.225)	(0.279)	(0.217)	(0.231)	(0.143)	(0.409)
Constant	36.60***	39.01***	37.29***	37.57***	36.54***	36.57***
	(0.603)	(0.789)	(0.593)	(0.649)	(0.395)	(1.154)
Observations	7,369	4,110	7,774	6,893	16,123	2,191
R-squared	0.192	0.162	0.200	0.208	0.178	0.173

Source:

Irish Labour Force Survey 2017 and 2018.

Notes:

The DiD model compares Q1/Q2 2017 with Q1/Q2 2018. Standard errors in parentheses with significance levels \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

When comparing the region of work specification (Table 9) to the region of residence specification (Table 10), note that the number of observations for Dublin in Table 10 is lower compared to Table 9. The opposite is the case for the other regions. This captures the fact that a lot of workers live outside Dublin but travel in to Dublin to work. If commuting workers are in some way different to individuals who live and work in the same location, then we may expect to see slight differences in the results from both specifications. For example, the hours effect in Dublin is greater in the region of residence specification compared to the region of work. The region of residence specification will not include the large number of workers from other regions who commute to Dublin to work. Given that commuters may invest a lot of time and money in travelling to work, they may be less likely to tolerate cuts to their hours compared to workers who live close to

their job. These types of differences between workers with different regions of work and residence and workers who live and work in the same region, may explain why the results in Tables 9 and 10 are slightly different.

TABLE 10 DIFFERENCE-IN-DIFFERENCES ESTIMATES OF THE EFFECT OF MINIMUM WAGE CHANGE ON HOURS WORKED BY REGION OF RESIDENCE

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	Border/ Mid-East	West	Mid-West/ South East	South West	Dublin	Midland
DiD	0.318	-2.351**	1.326*	1.446*	-2.052***	-2.544*
	(0.701)	(0.991)	(0.695)	(0.803)	(0.668)	(1.317)
Time	0.291	0.186	0.138	0.171	0.213	0.674*
	(0.204)	(0.285)	(0.227)	(0.241)	(0.164)	(0.377)
Minimum Wage	-9.067***	-6.465***	-9.187***	-8.848***	-8.995***	-8.509***
	(0.531)	(0.743)	(0.509)	(0.581)	(0.444)	(0.975)
Age	-0.0200**	-0.0632***	-0.0255***	-0.0380***	-0.0219***	-0.0520***
	(0.00854)	(0.0119)	(0.00931)	(0.00995)	(0.00668)	(0.0162)
Female	-7.257***	-5.895***	-7.696***	-8.238***	-5.770***	-6.539***
	(0.199)	(0.277)	(0.217)	(0.233)	(0.156)	(0.369)
Children (binary)	-0.499**	-0.299	-0.0852	-0.819***	-0.798***	0.464
	(0.200)	(0.276)	(0.218)	(0.236)	(0.161)	(0.370)
Medium ISCED (ref = low ISCED)	2.477***	1.693***	1.924***	2.218***	2.078***	0.691
	(0.318)	(0.490)	(0.340)	(0.396)	(0.285)	(0.580)
High ISCED	4.303***	3.739***	3.508***	4.519***	5.828***	2.056***
	(0.320)	(0.498)	(0.349)	(0.394)	(0.275)	(0.606)
2 <sup>nd</sup> Quarter	-0.215	0.689**	0.206	-0.0104	0.0167	0.145
	(0.195)	(0.271)	(0.213)	(0.230)	(0.156)	(0.360)
Constant	36.78***	37.98***	37.68***	37.80***	36.07***	39.13***
	(0.536)	(0.776)	(0.585)	(0.646)	(0.423)	(1.009)
Observations	9,626	4,387	8,212	6,934	13,749	2,612
R-squared	0.187	0.166	0.197	0.213	0.193	0.185

Source: Irish Labour Force Survey 2017 and 2018.

Notes: The DiD model compares Q1/Q2 2017 with Q1/Q2 2018. Standard errors in parentheses with significance levels \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

When employing a DiD estimator, it is good practice to estimate the model on a time period during which no policy change occurred, a so-called 'placebo' time period. If significant effects were detected in the placebo time period, this may call into question the validity of the DiD estimator, as this may indicate that there were diverging trends between the treatment and control groups even before the policy change occurred; any statistically significant impact observed for the treatment period could simply be a product of such ongoing divergence. To verify the validity

of our results and confirm that they are not being driven by diverging trends, we estimate our models using Quarters 1 and 2 of 2017 and Quarters 3 and 4 of 2017, during which time no minimum wage change occurred. Our ability to employ additional placebo tests on earlier time periods is constrained for two reasons. Firstly, the minimum wage question was added for the first time to the LFS in Quarter 2 of 2016. Secondly, we cannot use the period 2016 to 2017 as a placebo test, due to the minimum wage change which occurred in January 2017. Nevertheless, to the extent that trend effects will tend to be continuous over time, estimating our models for separate periods within a particular year during which the minimum wage was constant, is a reasonable robustness test.

The placebo results are shown in Tables 11 and 12. We observe no statistically significant results in the Sector model (Table 11). Likewise, in the Region model (Table 12), we observe no statistically significant negative hours effect. There is a positive hours effect for the Midland region, however this is only marginally significant. This implies that our results for the treatment period are unlikely to be attributable to diverging trends. Therefore, we can rule this out as a factor explaining the observed changes in hours worked. Furthermore, given that our DiD models explicitly control for seasonality, we can also exclude this as a factor. The exclusion of both seasonal and trend explanations leads greater weight to the likelihood that the observed impacts were related to the 2018 minimum wage increase.

The placebo for the Midland was positive at a 10 per cent level of confidence, however, in addition to being only weakly significant, no substantial or consistent impacts were detected for this region following the 2018 rate rise.

TABLE 11 DIFFERENCE-IN-DIFFERENCES PLACEBO ESTIMATES BY SECTOR

	(1)	(2)	(3)	(4)
VARIABLES	Wholesale and Retail	Accommodation and Food	Manufacturing	Other
DiD	0.0500	0.164	-0.772	0.647
	(0.663)	(0.834)	(0.821)	(0.517)
Time	-0.940***	-1.254**	0.658**	0.243
	(0.359)	(0.580)	(0.264)	(0.156)
Minimum Wage	-9.158***	-8.054***	-2.689***	-7.546***
	(0.478)	(0.599)	(0.599)	(0.366)
Age	0.0826***	0.135***	-0.0143	-0.0923***
	(0.0101)	(0.0161)	(0.00873)	(0.00478)
Female	-7.036***	-4.537***	-4.328***	-7.184***
	(0.244)	(0.380)	(0.197)	(0.112)
Children (binary)	-0.102	-0.935**	-0.637***	-1.032***
	(0.246)	(0.383)	(0.185)	(0.112)
Medium ISCED (ref = low ISCED)	3.198***	4.174***	0.802***	1.677***
	(0.363)	(0.547)	(0.299)	(0.198)
High ISCED	6.106***	8.629***	2.392***	3.683***
	(0.398)	(0.599)	(0.300)	(0.193)
2 <sup>nd</sup> Quarter	-0.00611	0.447	0.0572	0.263*
	(0.327)	(0.504)	(0.247)	(0.145)
3 <sup>rd</sup> Quarter	0.963***	2.149***	0.252	0.356**
	(0.361)	(0.559)	(0.274)	(0.166)
Constant	31.04***	25.78***	39.29***	40.57***
	(0.625)	(0.949)	(0.528)	(0.313)
Observations	6,560	3,276	5,458	29,287
R-squared	0.249	0.248	0.104	0.168

Source: Irish Labour Force Survey 2017 and 2018.

Notes: The DiD model compares Q1/Q2 2017 with Q3/Q4 2017. Standard errors in parentheses with significance levels \*\*\*

p<0.01, \*\* p<0.05, \* p<0.1.

TABLE 12 DIFFERENCE-IN-DIFFERENCES PLACEBO ESTIMATES BY REGION

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	Border/ Mid-East	West	Mid-West/ South East	South West	Dublin	Midland
DiD	-0.1000	-0.834	-0.617	-0.153	-0.605	2.422*
	(0.772)	(1.046)	(0.710)	(0.789)	(0.620)	(1.398)
Time	0.233	1.012**	-0.0196	-0.245	-0.0279	1.750***
	(0.329)	(0.412)	(0.316)	(0.342)	(0.207)	(0.606)
Minimum Wage	-8.983***	-7.121***	-8.659***	-8.724***	-8.858***	-8.594***
	(0.572)	(0.782)	(0.516)	(0.566)	(0.429)	(0.997)
Age	-0.0287***	-0.0713***	-0.0250***	-0.0347***	-0.0184***	0.0181
	(0.00981)	(0.0126)	(0.00943)	(0.0102)	(0.00624)	(0.0182)
Female	-7.266***	-5.490***	-7.614***	-8.430***	-6.028***	-6.073***
	(0.236)	(0.293)	(0.223)	(0.237)	(0.143)	(0.423)
Children (binary)	-0.480**	-0.901***	-0.0183	-1.021***	-0.955***	0.678
	(0.237)	(0.293)	(0.223)	(0.240)	(0.146)	(0.425)
Medium ISCED (ref = low ISCED)	2.500***	0.778	2.026***	2.199***	1.947***	1.879***
	(0.352)	(0.508)	(0.347)	(0.401)	(0.265)	(0.664)
High ISCED	4.100***	2.392***	3.779***	4.187***	5.318***	3.214***
	(0.363)	(0.513)	(0.357)	(0.399)	(0.256)	(0.699)
Quarter 2	0.0736	0.728*	0.176	0.0507	0.155	0.843
	(0.319)	(0.389)	(0.296)	(0.317)	(0.187)	(0.555)
Quarter 3	0.392	0.349	0.584*	1.060***	0.712***	-0.591
	(0.335)	(0.424)	(0.320)	(0.347)	(0.221)	(0.614)
Constant	36.53***	39.30***	37.21***	37.77***	36.57***	33.93***
	(0.625)	(0.822)	(0.600)	(0.665)	(0.399)	(1.160)
Observations	7,072	4,132	7,887	6,583	15,876	2,037
R-squared	0.188	0.139	0.201	0.226	0.185	0.157

Source:

Irish Labour Force Survey 2017 and 2018.

Notes:

The DiD model compares Q1/Q2 2017 with Q3/Q4 2017. Standard errors in parentheses with significance levels \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Region of work used in the analysis.

While we observe changes in hours worked in some regions and in the manufacturing sector, the question of what exactly is driving these movements remains open. In theory, changes to hours following the rate rise could be due to behavioural responses of either employers or employees. Employers may respond by reducing the hours of existing employees, thereby explaining the hours reduction. An alternative explanation may be due to compositional effects of labour supply, if the minimum wage entices more part-time minimum wage

This could also be due to a combination of both employer and employee behavioural responses.

employees into the labour market. To disentangle these two effects, we restrict our sample to incumbent employees who were with the same employer prior to the 2018 rate rise. If we observe that the impacts in hours worked were experienced by the incumbent sample, this would provide evidence against compositional effects.

In Tables 13 and 14, we re-estimate our DiD model on hours worked restricting our sample to individuals who began their current job before, or during, 2017. This ensures that only individuals who were in their job both before, and after, the 2018 rate rise are included in the data, thereby allowing us to measure the impact of the policy on the hours worked of incumbent employees only. The results are largely unchanged from those in Tables 8 and 9. Therefore, the changes in hours worked observed in our earlier models relate to impacts felt by existing employees, and not to compositional changes driven by flows of employees into, or out of, regions or sectors.

We next undertake a series of robustness tests, using our sample of incumbent employees. These include (i) dropping public sector employees from the sample on the grounds that many of these will also have received a pay increase in January 2018 and given the fact that MW employees do not typically work in the public sector, 11 (ii) adding a marital dummy to the models, and (iii) adding firm size and occupational controls to the models. We also test for the persistence of impacts beyond the first quarter of 2018.

On January 1, 2018, public sector annualised salaries increased by 1 per cent as part of the Public Service Stability Agreement.

TABLE 13 DIFFERENCE-IN-DIFFERENCES HOURS ESTIMATES BY SECTOR (CONDITIONAL ON STARTING THE JOB BEFORE 2018)

	(1)	(2)	(3)	(4)
VARIABLES	Wholesale and Retail	Accommodation and Food	Manufacturing	Other
DiD	-0.715	0.199	-1.644**	0.826
	(0.697)	(0.869)	(0.822)	(0.536)
Time	-0.578**	-0.489	0.768***	0.406***
	(0.277)	(0.463)	(0.186)	(0.115)
Minimum Wage	-9.121***	-8.275***	-2.900***	-7.493***
	(0.486)	(0.609)	(0.578)	(0.370)
Age	0.0834***	0.121***	-0.0223***	-0.101***
	(0.0106)	(0.0164)	(0.00851)	(0.00490)
Female	-6.889***	-4.040***	-4.094***	-7.299***
	(0.253)	(0.393)	(0.193)	(0.113)
Children (binary)	0.0757	-0.414	-0.209	-0.924***
	(0.255)	(0.397)	(0.181)	(0.113)
Medium ISCED (ref =low ISCED)	3.582***	3.863***	0.592**	1.630***
	(0.375)	(0.573)	(0.296)	(0.202)
High ISCED	6.632***	8.129***	2.026***	3.689***
	(0.410)	(0.620)	(0.297)	(0.197)
Quarter 2	-0.0283	0.380	0.188	0.110
	(0.251)	(0.388)	(0.179)	(0.111)
Constant	30.48***	26.07***	39.50***	41.03***
	(0.639)	(0.960)	(0.510)	(0.316)
Observations	6,224	3,070	5,230	28,379
R-squared	0.252	0.228	0.106	0.171

Source: Irish Labour Force Survey 2017 and 2018.

Notes: The DiD model compares Q1/Q2 2017 with Q1/Q2 2018. Standard errors in parentheses with significance levels \*\*\* p<0.01,

\*\* p<0.05, \* p<0.1.

TABLE 14 DIFFERENCE-IN-DIFFERENCES HOURS ESTIMATES BY REGION (CONDITIONAL ON STARTING THE JOB BEFORE 2018)

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	Border/ Mid-East	West	Mid-West/ South East	South West	Dublin	Midland
DiD	-0.627	-1.939*	0.817	1.160	-1.576**	-1.563
	(0.776)	(1.044)	(0.728)	(0.828)	(0.681)	(1.456)
Time	0.472*	0.325	0.132	0.421*	0.412***	0.804*
	(0.243)	(0.299)	(0.236)	(0.249)	(0.155)	(0.438)
Minimum Wage	-9.085***	-6.680***	-8.747***	-8.977***	-8.922***	-8.875***
	(0.571)	(0.772)	(0.511)	(0.578)	(0.439)	(1.041)
Age	-0.0325***	-0.0771***	-0.0273***	-0.0486***	-0.0241***	-0.0320*
	(0.00994)	(0.0125)	(0.00969)	(0.0104)	(0.00647)	(0.0185)
Female	-7.328***	-5.615***	-7.732***	-8.129***	-5.999***	-6.441***
	(0.236)	(0.291)	(0.226)	(0.240)	(0.149)	(0.427)
Children (binary)	-0.495**	-0.647**	-0.125	-0.783***	-0.619***	0.435
	(0.238)	(0.289)	(0.227)	(0.243)	(0.151)	(0.429)
Medium ISCED (ref =low ISCED)	2.544***	1.468***	2.108***	2.036***	1.853***	0.866
	(0.355)	(0.505)	(0.354)	(0.414)	(0.274)	(0.682)
High ISCED	4.045***	3.336***	3.732***	4.359***	5.176***	2.416***
	(0.364)	(0.513)	(0.363)	(0.412)	(0.264)	(0.718)
Quarter 2	-0.122	0.340	0.182	0.102	0.0920	0.609
	(0.231)	(0.284)	(0.221)	(0.237)	(0.148)	(0.417)
Constant	36.85***	38.93***	37.42***	38.04***	36.73***	37.27***
	(0.618)	(0.804)	(0.607)	(0.674)	(0.408)	(1.180)
Observations	6,942	3,947	7,425	6,414	15,159	2,098
R-squared	0.197	0.157	0.201	0.216	0.177	0.174

Source:

Irish Labour Force Survey 2017 and 2018.

Notes:

The DiD model compares 2017 Q1, Q2 with 2018 Q1, Q2. Standard errors in parentheses with significance levels \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Region of work used in the analysis.

The results for the sectoral models are reported in Table A1 in the Appendix. The negative hours effect for manufacturing is robust to all specifications. In terms of the persistence of the sectoral impacts, when we estimate the model for Q2/Q3 2017 to Q2/Q3 2018, and Q3/Q4 2017 to Q3/Q4 2018, the coefficients are broadly similar. However, the impacts for the Q2/Q3 and Q3/Q4 models are either not statistically significant or only marginally significant. Therefore, we cannot be fully confident that the fall in hours worked among minimum wage employees in the manufacturing sector persisted into the second half of 2018.

Table A2 shows the results of the robustness tests for the regional models. The estimated impacts for Dublin and the West are robust to the alternative specifications. With respect to the West, the estimated hours effect is substantially

larger when controls for occupation and firm size were added to the model. In terms of persistence, we detect no statistically significant regional hours effects for the Q3/Q4 models. Therefore, the fall in hours worked among minimum wage employees may not have persisted beyond the first six months of 2018. The lack of any strong evidence related to persistent impacts at the regional model is, perhaps, not surprising given the macroeconomic context of strong economic growth, falling unemployment and rising demand for labour. However, the fact that such regional variations existed at all leave open the possibility that the duration of adjustments could become more extensive in depressed labour markets.

## **SECTION 5**

## **Summary and conclusion**

This study tests for impacts in the hours worked of MW employees following the 2018 rate change, which saw the minimum wage increase from €9.25 to €9.55 per hour. In addition to analysing the impact of the rate change at a national level, we also test for the possibility of heterogeneous impacts occurring at the level of both region and sector. The rationale for the approach is to investigate whether findings of impacts carried out at a national level mask heterogeneous impacts that are restricted to minimum wage employees located in particular regions or sectors.

Approximately 8 per cent of all workers in Ireland were in minimum wage employment in 2017 and 2018. However, our analysis has shown that the incidence varies substantially by area, from 5 per cent of total employees in Dublin to over 10 per cent in the Border/Mid-East, Mid-West/South East and the Midland regions. With respect to sector, 30 per cent of all employees in the accommodation and food sector earn the minimum wage or less. In wholesale and retail, the incidence is approximately 16 per cent. These two sectors alone account for approximately 55 per cent of all minimum wage employees in Ireland. The incidence of minimum wage employment among manufacturing employees is lower, at 5.4 per cent, with these workers accounting for 8 per cent of all minimum wage employees.

In addition to variations in the incidence, we also found heterogeneity in the typical profile of minimum wage employees across regions and sectors. Minimum wage employees in Dublin tend to be younger, better educated, work on temporary contracts and have lower tenure than most other regions. With regard to sector, the profile of minimum wage workers in manufacturing looks very different to minimum wage employees in other sectors. Minimum wage employees in the manufacturing sector are more likely to be older, work full-time and have permanent jobs. For such employees, minimum wage employment is likely to represent their career. Therefore, based on their profile, these employees are likely to be reliant on statutory minimum wage increases to boost their incomes.

Our results show that the 2018 MW increase did not lead to any changes in the overall average hours worked of minimum wage employees, when looking at the full national sample. However, the research also demonstrates that such aggregate analysis can mask heterogeneous impacts that can occur at the regional or sectoral level. With respect to region, in the six months following the 2018 rate rise, average hours worked by minimum wage employees, relative to non-minimum wage employees, fell by 1.6 hours per week in Dublin and 1.7 hours per week in the West

region. The sectoral models indicate that the average hours worked of minimum wage employees, relative to non-minimum wage employees, in the manufacturing sector fell by 1.6 hours per week in the six-month period following the rate rise. In order to ensure that the estimated impacts were not caused by ongoing diverging trends in the hours worked of minimum and non-minimum wage workers, we reestimated our models for the period Q1/Q2 2017 to Q3/Q4 2017, during which the MW was constant. If the observed results were due to diverging pre-treatment trends, then it is reasonable to expect these results to be evident during this 2017 placebo period. However, no negative hours impacts were detected at either the regional or sectoral level in the placebo period. Therefore, this suggests that the hours effects observed in the first half of 2018 were linked to the minimum wage increase, as opposed to diverging trends between high and low paid workers.

Our central results were robust to a series of alternative specifications, and our regional results were broadly consistent when we used region of residence as opposed to region of work. We also show that the impacts were felt principally by minimum wage employees who were in their job prior to the 2018 rate change. Therefore, the hours effects were not driven by compositional changes related to higher flows of part-time workers into certain sectors or regions following the minimum wage increase. However, our analysis suggests that the observed impacts for Dublin, the West and the manufacturing sector did not appear to persist into the second half of 2018. It is possible that the immediate negative impact on the hours of minimum wage employees in the West, Dublin, and in the manufacturing sector would have been subsequently eased because of the continued tightening of the Irish labour market throughout 2018. However, the finding of possible heterogeneous impacts arising from the minimum wage change suggests that the existence of such variations should be monitored, particularly given that any impacts could become more persistent during periods of lower economic growth.

While the results from the study point towards a causal influence between the 2018 rate change and the heterogeneous change in hours worked of MW employees, some caution is required both in terms of the interpretation of results and the potential policy implications. While our placebo tests indicate that our results are not due to diverging pre-treatment trends in hours worked between MW and non-MW workers, the number of placebo periods available to us is limited, meaning that we cannot definitively rule out diverging trends as a factor. Secondly, despite the fact that our analysis points towards the changes being felt predominantly by incumbent MW workers, we cannot be certain about the exact mechanisms that brought about these changes and, consequently, the implications for policy. It could be that some of the observed impacts were due to employers reducing the hours of MW employees following the rate rise in order to reduce labour costs. However, it could also be the case that employees chose to substitute

more leisure for less work as a result of the rise in pay. The data at hand do not allow us to identify the potential behavioural responses that brought about the observed falls in hours worked. There is a need for greater insights into exactly how firms and employees behave following MW rate changes, so that policy can be properly informed. This is particularly the case in instances, such as those in the manufacturing sector, where impacts are observed among individuals who appear to be career MW employees.

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## **APPENDIX**

TABLE A1 ROBUSTNESS CHECKS FOR SECTOR

	Wholesale and Retail	Accommodatio n and Food	Manufacturing	Other
Baseline (conditional on starting job prior to 2018)	-0.715	0.199	-1.644**	0.826
	(0.697)	(0.869)	(0.822)	(0.536)
Robustness Checks				
Drop public sector employees (N/A for first 3 sectors as are mutually exclusive)	-0.715	0.199	-1.644**	0.365
	(0.697)	(0.869)	(0.822)	(0.553)
Add married dummy	-0.762	0.262	-1.644**	0.821
	(0.696)	(0.867)	(0.822)	(0.536)
Add firm size dummies and 1- digit occupation dummies	-1.358**	0.336	-2.243***	0.329
	(0.691)	(0.872)	(0.815)	(0.556)
Compare Quarters 2 and 3 of 2017 and 2018	0.314	-0.800	-1.439	0.239
	(0.766)	(0.944)	(0.930)	(0.588)
Compare Quarters 3 and 4 of 2017 and 2018	0.732	-0.132	-1.781*	1.675***
	(0.799)	(1.024)	(0.994)	(0.644)

Source: Irish Labour Force Survey 2017 and 2018.

Notes:

All robustness checks conditional on starting the job prior to 2018. Public sector employees refer to NACE sector 84 corresponding to 'Public administration and defence; compulsory social security'. Firm size takes on many categories denoting exact number of persons between one and ten, 11 to 19 persons, 20 to 49 persons, 50 to 99 persons, 100 to 249 persons, 250 to 500 persons, more than 500 persons. Ten one-digit occupation dummies were also added to the regression.

TABLE A2 ROBUSTNESS CHECKS FOR REGION

	Border/ Mid-East	West	Mid- West/ South East	South West	Dublin	Midland
Baseline (conditional on	-0.627	-1.939*	0.817	1.160	-1.576**	-1.563
starting job prior to 2018)	(0.776)	(1.044)	(0.728)	(0.828)	(0.681)	(1.456)
<b>Robustness Checks</b>						
Drop public sector	-0.968	-1.913*	0.567	1.105	-1.732**	-1.834
employees	(0.789)	(1.061)	(0.744)	(0.843)	(0.691)	(1.496)
Add married dummy	-0.662 (0.776)	-1.922* (1.044)	0.836 (0.728)	1.157 (0.828)	-1.577** (0.681)	-1.549 (1.455)
Add firm size dummies and 1- digit occupation dummies	-1.592** (0.785)	-3.114*** (1.068)	0.318 (0.725)	0.565 (0.843)	-1.583** (0.700)	-2.033** (1.488)
Compare Quarters 2 and 3 of 2017 and 2018	-0.980 (0.849)	-0.746 (1.160)	0.869 (0.802)	0.845 (0.897)	-1.664** (0.776)	-0.411 (1.473)
Compare Quarters 3 and 4	1.335	-0.340	1.560*	2.471**	0.185	-1.104
of 2017 and 2018	(0.880)	(1.254)	(0.875)	(0.973)	(0.845)	(1.571)

Source: Irish Labour Force Survey 2017 and 2018.

Notes:

All robustness checks conditional on starting the job prior to 2018. Region refers to region of work. Public sector employees refer to NACE sector 84 corresponding to 'Public administration and defence; compulsory social security'. Firm size takes on many categories denoting exact number of persons between one and ten 11 to 19 persons, 20 to 49 persons, 50 to 99 persons, 100 to 249 persons, 250 to 500 persons, more than 500 persons. Ten one-digit occupation dummies were also added to the regression.

TABLE A3 COUNTIES WITHIN NUTS 3 REGIONS IN 2018

NUTS 3 REGION	Counties	NUTS 3 REGION	Counties
Border	Donegal	South West	Cork
	Sligo		Kerry
	Leitrim		
	Cavan	Dublin	Dublin
	Monaghan		
		Mid-East	Wicklow
West	Galway		Kildare
	Mayo		Meath
	Roscommon		Louth
Mid-West	Clare	Midland	Longford
	Tipperary		Westmeath
	Limerick		Offaly
South East	Waterford		Laois
	Kilkenny		
	Carlow		
	Wexford		

Source: Irish Labour Force Survey 2017 and 2018.

TABLE A4 REGION CLASSIFICATION USED IN THE REPORT

NUTS 3 REGION	Counties	NUTS 3 REGION	Counties
Border/Mid-East	Donegal	South East/Mid-West	Clare
	Sligo		Tipperary
	Leitrim		Limerick
	Cavan		Waterford
	Monaghan		Kilkenny
	Wicklow		Carlow
	Kildare		Wexford
	Meath		
	Louth	South West	Cork
			Kerry
West	Galway		
	Mayo	Dublin	Dublin
	Roscommon		
		Midland	Longford
			Westmeath
			Offaly
			Laois

Source: Irish Labour Force Survey 2017 and 2018.

Notes: In Quarter 1, 2018, the NUTS3 boundaries were amended slightly such that county Louth changed from being included in the Border region to being included in the Mid-East region and Tipperary changed from the South East region to the Mid-West region. Therefore, we group the Border and Mid-East regions and South East and Mid-West regions.

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